

2018 秋季本科时间序列  
第四次作业答案

2018 年 11 月 13 日

1. (a)  $\{\theta_i\}_{i=0}^{\infty}$  绝对和收敛从而有部分和序列  $T_k = \sum_{i=0}^k |\theta_i|$  为 Cauchy 序列。  
对  $\forall \frac{\epsilon}{\bar{a}} > 0$ ,  $\exists K \in \mathbb{N}$ , 使得  $\forall m > n \geq K$ , 有

$$|T_m - T_n| = \sum_{i=n+1}^m |\theta_i| \leq \frac{\epsilon}{\bar{a}}$$

从而, 对  $S_k = \sum_{i=0}^k a_i \theta_i$ , 且  $|a_i| \leq \bar{a}$ , 有

$$|S_m - S_n| = \left| \sum_{i=n+1}^m a_i \theta_i \right| \leq \sum_{i=n+1}^m |a_i \theta_i| \leq \bar{a} \sum_{i=n+1}^m |\theta_i| \leq \bar{a} \frac{\epsilon}{\bar{a}} = \epsilon$$

因此, 部分和序列  $S_k$  为 Cauchy 序列, 所以级数  $\sum_{i=0}^{\infty} a_i \theta_i$  收敛

- (b) 因为  $\{Y_i\}_{i=0}^{\infty}$  为平稳序列

对  $\forall t$ , 有  $\mathbb{E}Y_t = \mu_Y$

且  $\mu_Y$  与  $t$  无关。从而有

$$\begin{aligned} \mathbb{E}X_t &= \mathbb{E} \left[ \sum_{i=0}^{\infty} \theta_i Y_{t-i} \right] = \sum_{i=0}^{\infty} \mathbb{E}[\theta_i Y_{t-i}] \\ &= \sum_{i=0}^{\infty} \theta_i \mathbb{E}[Y_{t-i}] \\ &= \mu_Y \sum_{i=0}^{\infty} \theta_i \end{aligned}$$

$\theta_i$  与  $t$  无关, 所以  $\mathbb{E}X_t$  与  $t$  无关

(c)

$$\begin{aligned}\operatorname{cov}(X_t, Y_{t-j}) &= \operatorname{cov}\left(\sum_{i=0}^{\infty} \theta_i Y_{t-i}, Y_{t-j}\right) \\ &= \sum_{i=0}^{\infty} \theta_i \operatorname{cov}(Y_{t-i}, Y_{t-j}) \\ &= \sum_{i=0}^{\infty} \theta_i \sigma_Y^2(j-i)\end{aligned}$$

$\theta_i$  和  $\sigma_Y^2(j-i)$  与  $t$  无关, 所以  $\operatorname{cov}(X_t, Y_{t-j})$  与  $t$  无关  
对  $\forall k, |\sigma_Y^2(k)| \leq \sigma_Y^2$ , 所以序列  $\{\sigma_Y^2(j-i)\}_{i=0}^{\infty}$  有界  
由题 (a) 结论, 级数

$$\operatorname{cov}(X_t, Y_{t-j}) = \sum_{i=0}^{\infty} \sigma_Y^2(j-i)\theta_i$$

收敛

(d)

$$\begin{aligned}\operatorname{cov}(X_t, X_{t-k}) &= \operatorname{cov}\left(X_t, \sum_{i=0}^{\infty} \theta_i Y_{t-k-i}\right) \\ &= \sum_{i=0}^{\infty} \theta_i \operatorname{cov}(X_t, Y_{t-k-i}) \\ &= \sum_{i=0}^{\infty} \theta_i \left[ \sum_{j=0}^{\infty} \sigma_Y^2(k+i-j)\theta_j \right] \\ &= \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \theta_i \theta_j \sigma_Y^2(k+i-j).\end{aligned}$$

同理,  $\theta_i, \theta_j$  和  $\sigma_Y^2(k+i-j)$  均与  $t$  无关

所以  $\operatorname{cov}(X_t, X_{t-k})$  与  $t$  无关。

由上述推导, 有  $\sum_{j=0}^{\infty} \theta_j \sigma_Y^2(k+i-j)$  有界,  $\{\theta_i\}_{i=0}^{\infty}$  收敛, 所以级数

$$\operatorname{cov}(X_t, X_{t-k}) = \sum_{i=0}^{\infty} \theta_i \sum_{j=0}^{\infty} \theta_j \sigma_Y^2(k+i-j)$$

收敛

2. 附加题

(a)  $\lambda, \eta$  为  $A(z) = 1 - \phi_1 z - \phi_2 z^2$  的零点, 代入有

$$\begin{cases} 1 - \phi_1 \lambda - \phi_2 \lambda^2 = 0 \\ 1 - \phi_1 \eta - \phi_2 \eta^2 = 0 \end{cases}$$

得  $\phi_1 = \frac{\lambda + \eta}{\lambda \eta}$ ,  $\phi_2 = -\frac{1}{\lambda \eta}$ , 从而有

$$A^{-1}(z) = \frac{1}{1 - \phi_1 z - \phi_2 z^2} = \frac{1}{1 - \frac{z}{\lambda}} \cdot \frac{1}{1 - \frac{z}{\eta}}$$

若

$$A^{-1}(z) = \frac{a}{1 - \frac{z}{\lambda}} + \frac{b}{1 - \frac{z}{\eta}}$$

对应可解

$$a = \frac{1}{1 - \frac{\lambda}{\eta}}, \quad b = \frac{1}{1 - \frac{\eta}{\lambda}}$$

进而

$$\begin{aligned} X_t &= A^{-1}(\mathcal{L})\varepsilon_t \\ &= \frac{1}{1 - \frac{\lambda}{\eta}} \cdot \frac{1}{1 - \frac{\mathcal{L}}{\lambda}} \varepsilon_t + \frac{1}{1 - \frac{\eta}{\lambda}} \cdot \frac{1}{1 - \frac{\mathcal{L}}{\eta}} \varepsilon_t \\ &= \frac{1}{1 - \frac{\lambda}{\eta}} \sum_{i=0}^{\infty} \left(\frac{1}{\lambda}\right)^i \mathcal{L}^i \varepsilon_t + \frac{1}{1 - \frac{\eta}{\lambda}} \sum_{j=0}^{\infty} \left(\frac{1}{\eta}\right)^j \mathcal{L}^j \varepsilon_t \\ &= \sum_{i=0}^{\infty} \left[ \frac{1}{1 - \frac{\lambda}{\eta}} \left(\frac{1}{\lambda}\right)^i + \frac{1}{1 - \frac{\eta}{\lambda}} \left(\frac{1}{\eta}\right)^i \right] \varepsilon_{t-i} \\ &= \sum_{i=0}^{\infty} \theta_i \varepsilon_{t-i} \end{aligned}$$

令  $\text{var}(\varepsilon_t) = \sigma^2$ , 进而计算  $X_t$  的方差

$$\begin{aligned}
\text{var}(X_t) &= \text{var}\left(\sum_{i=0}^{\infty} \theta_i \varepsilon_{t-i}\right) \\
&= \sigma^2 \sum_{i=0}^{\infty} \theta_i^2 \\
&= \sigma^2 \sum_{i=0}^{\infty} \left[ \frac{1}{1 - \frac{\lambda}{\eta}} \left(\frac{1}{\lambda}\right)^i + \frac{1}{1 - \frac{\eta}{\lambda}} \left(\frac{1}{\eta}\right)^i \right]^2 \\
&= \sigma^2 \left[ \left(\frac{\eta}{\eta - \lambda}\right)^2 \sum_{i=0}^{\infty} \left(\frac{1}{\lambda^2}\right)^i - \frac{2\lambda\eta}{(\eta - \lambda)^2} \sum_{i=0}^{\infty} \left(\frac{1}{\lambda\eta}\right)^i + \left(\frac{\lambda}{\eta - \lambda}\right)^2 \sum_{i=0}^{\infty} \left(\frac{1}{\eta^2}\right)^i \right] \\
&= \sigma^2 \left[ \left(\frac{\eta}{\eta - \lambda}\right)^2 \lim_{i \rightarrow \infty} \left(\frac{1 - \frac{1}{\lambda^{2i}}}{1 - \frac{1}{\lambda^2}}\right) - \frac{2\lambda\eta}{(\eta - \lambda)^2} \lim_{i \rightarrow \infty} \left(\frac{1 - \frac{1}{\lambda\eta^i}}{1 - \frac{1}{\lambda\eta}}\right) + \left(\frac{\lambda}{\eta - \lambda}\right)^2 \lim_{i \rightarrow \infty} \left(\frac{1 - \frac{1}{\eta^{2i}}}{1 - \frac{1}{\eta^2}}\right) \right] \\
&= \sigma^2 \left(\frac{\lambda\eta}{\eta - \lambda}\right)^2 \left[ \frac{1}{\lambda^2 - 1} - \frac{2}{\lambda\eta - 1} + \frac{1}{\eta^2 - 1} \right] \\
&= \sigma^2 \left(\frac{\lambda\eta}{\eta - \lambda}\right)^2 \left[ \frac{\lambda^3\eta + \lambda\eta^3 + \lambda^2 + \eta^2 - 2\lambda\eta - 2\lambda^2\eta^2}{(\lambda^2 - 1)(\lambda\eta - 1)(\eta^2 - 1)} \right] \\
&= \frac{\sigma^2(\lambda\eta + 1)(\lambda\eta)^2}{(\lambda\eta - 1)[\lambda^2\eta^2 - (\lambda^2 + \eta^2) + 1]}
\end{aligned}$$

(b) 根据  $X_t$  表达式, 有:

$$\begin{aligned}
\text{var}(X_t) &= \text{cov}(X_t, X_t) \\
&= \text{cov}\left(\frac{1}{(1 - \mathcal{L}/\lambda)} Y_t, \frac{1}{(1 - \mathcal{L}/\lambda)} Y_t\right) \\
&= \text{cov}\left(\sum_{i=0}^{\infty} \left(\frac{1}{\lambda}\right)^i Y_{t-i}, \sum_{j=0}^{\infty} \left(\frac{1}{\lambda}\right)^j Y_{t-j}\right) \\
&= \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \left(\frac{1}{\lambda}\right)^{i+j} \text{cov}(Y_{t-i}, Y_{t-j})
\end{aligned}$$

又  $Y_t = \frac{1}{(1 - \mathcal{L}/\lambda)} \varepsilon_t$ , 所以  $Y_t$  为 AR(1) 过程:

$$Y_t = \frac{1}{\lambda} Y_{t-1} + \varepsilon_t$$

所以有:

$$\sigma_Y(k) = \text{cov}(Y_t, Y_{t-k}) = \left(\frac{1}{\lambda}\right)^k \frac{1}{1 - (1/\lambda)^2} \sigma^2$$

所以有  $\sigma_Y^2(j-i) = \left(\frac{1}{\lambda}\right)^{j-i} \frac{1}{1-(1/\lambda)^2} \sigma^2$ , 代入得:

$$\begin{aligned}
 \text{var}(X_t) &= \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \left(\frac{1}{\lambda}\right)^{i+j} \text{cov}(Y_{t-i}, Y_{t-j}) \\
 &= \frac{1}{1-(1/\lambda)^2} \sigma^2 \sum_{i=0}^{\infty} \sum_{j=0}^{\infty} \left(\frac{1}{\lambda}\right)^{i+j} \left(\frac{1}{\lambda}\right)^{|j-i|} \\
 &= \frac{1}{1-(1/\lambda)^2} \sigma^2 \sum_{i=0}^{\infty} \left[ \sum_{j=0}^i \left(\frac{1}{\lambda}\right)^{2i} + \sum_{j=i+1}^{\infty} \left(\frac{1}{\lambda^2}\right)^j \right] \\
 &= \frac{1}{1-(1/\lambda)^2} \sigma^2 \left[ \sum_{i=0}^{\infty} \frac{i}{\lambda^{2i}} + \frac{\lambda^2}{(\lambda^2-1)^2} \right] \\
 &= \frac{1}{1-(1/\lambda)^2} \sigma^2 \left[ \frac{\lambda^4}{(\lambda^2-1)^2} + \frac{\lambda^2}{(\lambda^2-1)^2} \right] \\
 &= \frac{\lambda^4(\lambda^2+1)}{(\lambda^2-1)^3}
 \end{aligned}$$

(c) 当  $\lambda \neq \eta$

$$\text{var}(X_t) = \frac{\sigma^2(\lambda\eta+1)(\lambda\eta)^2}{(\lambda\eta-1)[\lambda^2\eta^2 - (\lambda^2+\eta)^2 + 1]} = \frac{\sigma^2(\lambda\eta+1)(\lambda\eta)^2}{(\lambda\eta-1)[\lambda^2\eta^2 - (\lambda+\eta)^2 + 2\lambda\eta + 1]}$$

因为  $\lambda, \eta$  是方程  $1 - \phi_1 z - \phi_2 z^2 = 0$  的两个解, 从而有:

$$\lambda + \eta = -\frac{\phi_1}{\phi_2}, \quad \lambda\eta = -\frac{1}{\phi_2}$$

代入原式得:

$$\text{var}(X_t) = \frac{(1-\phi_2)\sigma^2}{(\phi_2+1)(\phi_2^2 - 2\phi_2 - \phi_1^2 + 1)}$$

当  $\lambda = \eta$  时,  $\lambda = -\frac{2}{\phi_1}$  代入得:

$$\text{var}(X_t) = \frac{16(4 + \phi_1^2)}{(4 - \phi_1^2)^3}$$