

2023秋季本科时间序列

第1次作业答案

9月22日

1. 根据分布函数定义, $F(x)$ 可写为

$$F(x) = P(X \leq x) = \int_{-\infty}^x f(t)dt = \int_{-\infty}^0 f(t)dt + \int_0^x f(t)dt$$

取 $a, b \in [-\infty, +\infty]$, $x + \Delta x \in [a, b]$

$$\text{令 } g(x) = \int_0^x f(t)dt, \text{ 则 } g(x + \Delta x) - g(x) = \int_x^{x+\Delta x} f(t)dt$$

由 $f(x)$ 在 R 上可积, 得 $f(x)$ 在 R 上有界, 即 $|f(x)| \leq M$

$$\text{则 } |g(x + \Delta x) - g(x)| \leq \int_x^{x+\Delta x} |f(t)| dt \leq M \Delta x, \therefore \lim_{\Delta x \rightarrow 0} (g(x + \Delta x) - g(x)) = 0$$

$$\therefore \int_0^x f(t)dt \text{ 在 } [a, b] \text{ 上连续, } \therefore F(x) = \int_{-\infty}^x f(t)dt = \int_{-\infty}^0 f(t)dt + \int_0^x f(t)dt \text{ 连续}$$

2. (a) \because 协方差矩阵为半正定矩阵且为对称矩阵, 方差为非负数

$$\therefore ad - bc \geq 0, b = c, a > 0, b > 0$$

(b) i. 由题已知, $Var(X) = E(X^2) - [E(X)]^2$

$\because [X, Y]^T$ 服从2元联合正态分布

$$\therefore f(x, y) = \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{-\frac{1}{2(1-\rho^2)} \left[\frac{(x-\mu_x)^2}{\sigma_x^2} - 2\rho \frac{(x-\mu_x)(y-\mu_y)}{\sigma_x\sigma_y} + \frac{(y-\mu_y)^2}{\sigma_y^2} \right]}$$

则

$$\begin{aligned} E(x^2) &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x^2 f(x, y) dx dy \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \frac{x^2}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{-\frac{1}{2(1-\rho^2)} \left[\frac{(x-\mu_x)^2}{\sigma_x^2} - 2\rho \frac{(x-\mu_x)(y-\mu_y)}{\sigma_x\sigma_y} + \frac{(y-\mu_y)^2}{\sigma_y^2} \right]} dx dy \end{aligned}$$

上式指数部分可改写为

$$-\frac{1}{2}\left(\rho\frac{x-\mu_x}{\sigma_x\sqrt{1-\rho^2}}-\frac{y-\mu_y}{\sigma_y\sqrt{1-\rho^2}}\right)^2-\frac{(x-\mu_x)^2}{2\sigma_x^2}$$

则

$$\begin{aligned} E(X^2) &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \frac{x^2}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{\left[-\frac{1}{2}\left(\rho\frac{x-\mu_x}{\sigma_x\sqrt{1-\rho^2}}-\frac{y-\mu_y}{\sigma_y\sqrt{1-\rho^2}}\right)^2-\frac{(x-\mu_x)^2}{2\sigma_x^2}\right]} dx dy \\ &= \int_{-\infty}^{\infty} \frac{x^2}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{-\frac{(x-\mu_x)^2}{2\sigma_x^2}} dx \int_{-\infty}^{\infty} e^{\left[-\frac{1}{2}\left(\rho\frac{x-\mu_x}{\sigma_x\sqrt{1-\rho^2}}-\frac{y-\mu_y}{\sigma_y\sqrt{1-\rho^2}}\right)^2\right]} dy \end{aligned}$$

对积分变量Y作如下替换:

$$t = \left(\rho\frac{x-\mu_x}{\sigma_x\sqrt{1-\rho^2}}-\frac{y-\mu_y}{\sigma_y\sqrt{1-\rho^2}}\right)$$

则上式可化为:

$$\begin{aligned} E(x^2) &= \int_{-\infty}^{\infty} \frac{x^2}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{-\frac{(x-\mu_x)^2}{2\sigma_x^2}} dx \int_{-\infty}^{\infty} e^{-\frac{t^2}{2}} \sigma_y\sqrt{1-\rho^2} dt \\ &= \int_{-\infty}^{\infty} \frac{x^2}{\sqrt{2\pi}\sigma_x} e^{-\frac{(x-\mu_x)^2}{2\sigma_x^2}} dx \\ &= \frac{\sigma_x^2}{\sqrt{2\pi}} \int_{-\infty}^{\infty} z^2 e^{-\frac{z^2}{2}} dz \\ &= \frac{\sigma_x^2}{\sqrt{2\pi}} \times \sqrt{2\pi} \\ &= \sigma_x^2 = a \end{aligned}$$

所以 $Var(X) = E(x^2) - (E(X))^2 = a$

ii. 由题已知, $Var(Y) = E(Y^2) - [E(Y)]^2$

$\therefore [X, Y]^T$ 服从2元联合正态分布

$$\therefore f(x, y) = \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{-\frac{1}{2(1-\rho^2)}\left[\frac{(x-\mu_x)^2}{\sigma_x^2}-2\rho\frac{(x-\mu_x)(y-\mu_y)}{\sigma_x\sigma_y}+\frac{(y-\mu_y)^2}{\sigma_y^2}\right]}$$

则

$$\begin{aligned} E(Y^2) &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} y^2 f(x, y) dx dy \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \frac{y^2}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{-\frac{1}{2(1-\rho^2)}\left[\frac{(x-\mu_x)^2}{\sigma_x^2}-2\rho\frac{(x-\mu_x)(y-\mu_y)}{\sigma_x\sigma_y}+\frac{(y-\mu_y)^2}{\sigma_y^2}\right]} dx dy \end{aligned}$$

上式指数部分可改写为

$$-\frac{1}{2}\left(\rho\frac{y-\mu_y}{\sigma_y\sqrt{1-\rho^2}}-\frac{x-\mu_x}{\sigma_x\sqrt{1-\rho^2}}\right)^2-\frac{(y-\mu_y)^2}{2\sigma_y^2}$$

则

$$\begin{aligned} E(Y^2) &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \frac{y^2}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{\left[-\frac{1}{2}\left(\rho\frac{y-\mu_y}{\sigma_y\sqrt{1-\rho^2}} - \frac{x-\mu_x}{\sigma_x\sqrt{1-\rho^2}}\right)^2 - \frac{(y-\mu_y)^2}{2\sigma_y^2}\right]} dx dy \\ &= \int_{-\infty}^{\infty} \frac{y^2}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{-\frac{(y-\mu_y)^2}{2\sigma_y^2}} dy \int_{-\infty}^{\infty} e^{\left[-\frac{1}{2}\left(\rho\frac{y-\mu_y}{\sigma_y\sqrt{1-\rho^2}} - \frac{x-\mu_x}{\sigma_x\sqrt{1-\rho^2}}\right)^2\right]} dx \end{aligned}$$

对积分变量X作如下替换:

$$t = \left(\rho\frac{y-\mu_y}{\sigma_y\sqrt{1-\rho^2}} - \frac{x-\mu_x}{\sigma_x\sqrt{1-\rho^2}}\right)$$

则上式可化为:

$$\begin{aligned} E(Y^2) &= \int_{-\infty}^{\infty} \frac{y^2}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{-\frac{(y-\mu_y)^2}{2\sigma_y^2}} dy \int_{-\infty}^{\infty} e^{-\frac{t^2}{2}} \sigma_x\sqrt{1-\rho^2} dt \\ &= \int_{-\infty}^{\infty} \frac{y^2}{\sqrt{2\pi}\sigma_y} e^{-\frac{(y-\mu_y)^2}{2\sigma_y^2}} dy \\ &= \frac{\sigma_y^2}{\sqrt{2\pi}} \int_{-\infty}^{\infty} z^2 e^{-\frac{z^2}{2}} dz \\ &= \frac{\sigma_y^2}{\sqrt{2\pi}} \times \sqrt{2\pi} \\ &= \sigma_y^2 = d \end{aligned}$$

所以 $Var(Y) = E(Y^2) - (E(Y))^2 = d$

iii. $cov(x,y) = E(XY) - E(X)E(Y) = E(XY)$

$$\begin{aligned} E(XY) &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xyf(x,y) dx dy \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \frac{xy}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{-\frac{1}{2(1-\rho^2)}\left[\frac{(x-\mu_x)^2}{\sigma_x^2} - 2\rho\frac{(x-\mu_x)(y-\mu_y)}{\sigma_x\sigma_y} + \frac{(y-\mu_y)^2}{\sigma_y^2}\right]} dx dy \end{aligned}$$

上式指数部分可改写为:

$$\frac{(x-\mu_x)^2}{\sigma_x^2} - 2\rho\frac{(x-\mu_x)(y-\mu_y)}{\sigma_x\sigma_y} + \frac{(y-\mu_y)^2}{\sigma_y^2} = \left(\frac{x-\mu_x}{\sigma_x} - \rho\frac{(y-\mu_y)}{\sigma_y}\right)^2 + (\sqrt{1-\rho^2}\frac{y-\mu_y}{\sigma_y})^2$$

对积分变量X作如下替换:

$$\begin{aligned} u &= \frac{1}{\sqrt{1-\rho^2}}\left(\frac{x-\mu_x}{\sigma_x} - \rho\frac{(y-\mu_y)}{\sigma_y}\right) \\ v &= \frac{y-\mu_y}{\sigma_y} \end{aligned}$$

则

$$\begin{aligned}
\text{cov}(x, y) &= \frac{\sigma_x \sigma_y}{2\pi} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (uv\sqrt{1-\rho^2} + \rho v^2) e^{-\frac{1}{2}(u^2+v^2)} du dv \\
&= \frac{\sigma_x \sigma_y}{2\pi} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (uv\sqrt{1-\rho^2} e^{-\frac{1}{2}(u^2+v^2)} + \rho v^2 e^{-\frac{1}{2}(u^2+v^2)}) du dv \\
&= \frac{\sigma_x \sigma_y}{2\pi} \times 2\pi \times \rho = \rho \sigma_x \sigma_y \\
&= b = c
\end{aligned}$$

(c)

$$\begin{aligned}
f_x(x) &= \int_{-\infty}^{\infty} \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{\left[-\frac{1}{2}\left(\rho\frac{x-\mu_x}{\sigma_x\sqrt{1-\rho^2}} - \frac{y-\mu_y}{\sigma_y\sqrt{1-\rho^2}}\right)^2 - \frac{(x-\mu_x)^2}{2\sigma_x^2}\right]} dy \\
&= \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{-\frac{(x-\mu_x)^2}{2\sigma_x^2}} \int_{-\infty}^{\infty} e^{\left[-\frac{1}{2}\left(\rho\frac{x-\mu_x}{\sigma_x\sqrt{1-\rho^2}} - \frac{y-\mu_y}{\sigma_y\sqrt{1-\rho^2}}\right)^2\right]} dy \\
&= \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{-\frac{(x-\mu_x)^2}{2\sigma_x^2}} \int_{-\infty}^{\infty} e^{-\frac{t^2}{2}} \sigma_y\sqrt{1-\rho^2} dt \\
&= \frac{1}{\sqrt{2\pi}\sigma_x} e^{-\frac{(x-\mu_x)^2}{2\sigma_x^2}} \\
&= \frac{1}{\sqrt{2\pi a}} e^{-\frac{x^2}{2a}} \\
f_y(y) &= \int_{-\infty}^{\infty} \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{\left[-\frac{1}{2}\left(\rho\frac{y-\mu_y}{\sigma_y\sqrt{1-\rho^2}} - \frac{x-\mu_x}{\sigma_x\sqrt{1-\rho^2}}\right)^2 - \frac{(y-\mu_y)^2}{2\sigma_y^2}\right]} dx \\
&= \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{-\frac{(y-\mu_y)^2}{2\sigma_y^2}} \int_{-\infty}^{\infty} e^{\left[-\frac{1}{2}\left(\rho\frac{y-\mu_y}{\sigma_y\sqrt{1-\rho^2}} - \frac{x-\mu_x}{\sigma_x\sqrt{1-\rho^2}}\right)^2\right]} dx \\
&= \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} e^{-\frac{(y-\mu_y)^2}{2\sigma_y^2}} \int_{-\infty}^{\infty} e^{-\frac{t^2}{2}} \sigma_x\sqrt{1-\rho^2} dt \\
&= \frac{1}{\sqrt{2\pi}\sigma_y} e^{-\frac{(y-\mu_y)^2}{2\sigma_y^2}} \\
&= \frac{1}{\sqrt{2\pi d}} e^{-\frac{y^2}{2d}}
\end{aligned}$$

3. (a) $\because \sum_{i=1}^n \sum_{j=1}^n \frac{i+j}{C} = 1$

$\therefore C = \sum_{i=1}^n \sum_{j=1}^n (i+j) = \sum_{i=1}^n \sum_{j=1}^n i + \sum_{i=1}^n \sum_{j=1}^n j = \frac{n(n+1)}{2} \times n + \frac{n(n+1)}{2} \times n = n^2(n+1) = n^3 + n^2$

(b) $\because X, Y$ 在题目中具有对称性

$$\begin{aligned}
 \therefore E(X) = E(Y) &= \sum_{i=1}^n \sum_{j=1}^n \frac{i(i+j)}{n^2(n+1)} = \sum_{i=1}^n \sum_{j=1}^n \frac{j(i+j)}{n^2(n+1)} \\
 &= \frac{1}{n^2(n+1)} \sum_{i=1}^n \sum_{j=1}^n (i^2 + ij) \\
 &= \frac{1}{n^2(n+1)} \sum_{i=1}^n \sum_{j=1}^n i^2 + \frac{1}{n^2(n+1)} \sum_{i=1}^n \sum_{j=1}^n ij \\
 &= \frac{1}{n^2(n+1)} \times n \times \frac{n(n+1)(2n+1)}{6} + \frac{1}{n^2(n+1)} \times \left(\frac{n(n+1)}{2}\right)^2 \\
 &= \frac{(2n+1)}{6} + \frac{n+1}{4} \\
 &= \frac{7n+5}{12}
 \end{aligned}$$

$$\begin{aligned}
 E(X^2) = E(Y^2) &= \sum_{i=1}^n \sum_{j=1}^n \frac{i^2(i+j)}{n^2(n+1)} \\
 &= \frac{1}{n^2(n+1)} \sum_{i=1}^n \sum_{j=1}^n (i^3 + i^2j) \\
 &= \frac{1}{n^2(n+1)} \sum_{i=1}^n \sum_{j=1}^n i^3 + \frac{1}{n^2(n+1)} \sum_{i=1}^n \sum_{j=1}^n i^2j \\
 &= \frac{1}{n^2(n+1)} \times n \times \left(\frac{n(n+1)}{2}\right)^2 + \frac{1}{n^2(n+1)} \times n \times \frac{n(n+1)}{2} \times \frac{n(n+1)(2n+1)}{6} \\
 &= \frac{n(n+1)}{4} + \frac{(n+1)(2n+1)}{6} = \frac{(n+1)(5n+1)}{12}
 \end{aligned}$$

$$\text{Var}(X) = \text{Var}(Y) = E(X^2) - (E(X))^2 = \frac{(n+1)(5n+1)}{12} - \left(\frac{7n+5}{12}\right)^2 = \frac{(n-1)(11n+13)}{144}$$

$$E(XY) = \sum_{i=1}^n \sum_{j=1}^n \frac{ij(i+j)}{n^2(n+1)} = \frac{1}{n^2(n+1)} \sum_{i=1}^n \sum_{j=1}^n i^2j + ij^2 = \frac{(n+1)(2n+1)}{6}$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y) = \frac{(n+1)(2n+1)}{6} - \left(\frac{7n+5}{12}\right)^2 = \frac{-(n-1)^2}{144}$$

4. (a)

$$\sum_{i=1}^n \sum_{j=1}^n \sum_{k=1}^n \frac{i+j+k}{C} = 1$$

$$C = \sum_{i=1}^n \sum_{j=1}^n \sum_{k=1}^n (i+j+k) = \frac{3n^3(n+1)}{2}$$

(b)

$$\begin{aligned} E(XYZ) &= \frac{1}{C} \sum_{i=1}^n \sum_{j=1}^n \sum_{k=1}^n (i+j+k)ijk \\ &= \frac{1}{C} \times 3 \times \sum_{i=1}^n \sum_{j=1}^n \sum_{k=1}^n i^2jk \\ &= \frac{(n+1)^2(2n+1)}{12} \end{aligned}$$