

2020 秋季本科时间序列

## 第 7 次作业答案

12 月 16 日

1. (a) 令  $Z = (X_i - \mathbb{E}X_i) - \Sigma_{XY,i}\Sigma_{YY}^{-1}(Y - \mathbb{E}Y)$

易得:

i.  $\mathbb{E}Z = 0$

证明:

$$\begin{aligned}\mathbb{E}Z &= \mathbb{E}[(X_i - \mathbb{E}X_i) - \Sigma_{XY,i}\Sigma_{YY}^{-1}(Y - \mathbb{E}Y)] \\ &= \mathbb{E}X_i - \mathbb{E}X_i - \Sigma_{XY,i}\Sigma_{YY}^{-1}(\mathbb{E}Y - \mathbb{E}Y) \\ &= 0\end{aligned}$$

ii.  $\text{cov}(Z, Y^T) = \vec{0}$

证明:

$$\begin{aligned}\text{cov}(Z, Y^T) &= \mathbb{E}[Z(Y - \mu_Y)^T] \\ &= \mathbb{E}[(X_i - \mathbb{E}X_i)(Y - \mu_Y)^T - \Sigma_{XY,i}\Sigma_{YY}^{-1}(Y - \mathbb{E}Y)(Y - \mu_Y)^T] \\ &= \Sigma_{XY,i} - \Sigma_{XY,i}\Sigma_{YY}^{-1}\Sigma_{YY} \\ &= \vec{0}\end{aligned}$$

所以  $Z$  和  $Y$  互相独立对于任意关于  $X_i$  的预测函数  $f(Y)$

有

$$\begin{aligned}\mathbb{E}[X_i - f(Y)]^2 &= \mathbb{E}[X_i - \mathbb{E}(X_i|Y) + \mathbb{E}(X_i|Y) - f(Y)]^2 \\ &= \mathbb{E}[X_i - \mathbb{E}(X_i|Y)]^2 + 2\mathbb{E}[(X_i - \mathbb{E}(X_i|Y))(\mathbb{E}(X_i|Y) - f(Y))] + \mathbb{E}[\mathbb{E}(X_i|Y) - f(Y)]^2 \\ &= \mathbb{E}[X_i - \mathbb{E}(X_i|Y)]^2 + 2\mathbb{E}[\mathbb{E}[(X_i - \mathbb{E}(X_i|Y))(\mathbb{E}(X_i|Y) - f(Y))|Y]] \\ &\quad + \mathbb{E}[\mathbb{E}(X_i|Y) - f(Y)]^2 \\ &= \mathbb{E}[X_i - \mathbb{E}(X_i|Y)]^2 + \mathbb{E}[\mathbb{E}(X_i|Y) - f(Y)]^2 \\ &\geq \mathbb{E}[X_i - \mathbb{E}(X_i|Y)]^2 \quad \text{当且仅当 } f(Y) = \mathbb{E}(X_i|Y) \text{ 时等式成立}\end{aligned}$$

由此可知  $\mathbb{E}(X_i|Y)$  是  $X_i$  关于  $Y$  的最小均方误差标准下的最优预测

$$\text{则 } \mathbb{E}[X_i - Z|Y] = \mathbb{E}(X_i|Y) - \mathbb{E}(Z|Y) = \mathbb{E}(X_i|Y)$$

可知  $X_i - Z$  是  $X_i$  关于  $Y$  的最优预测

$$\text{则 } \mathbb{E}(X_i|Y) = \mathbb{E}[X_i - Z|Y] = \mathbb{E}[\mathbb{E}X_i - \Sigma_{XY,i}\Sigma_{YY}^{-1}(Y - \mu_Y)|Y]$$

$$= \mathbb{E}X_i + \Sigma_{XY,i}\Sigma_{YY}^{-1}(Y - \mu_Y) \text{ 得证}$$

(b) 由 (a) 问可知  $\mathbb{E}(X_i|Y) = \mathbb{E}X_i + \Sigma_{XY,i}\Sigma_{YY}^{-1}(Y - \mu_Y)$

$$\begin{aligned} \text{则 } \mathbb{E}(X|Y) &= \mathbb{E} \begin{pmatrix} X_1|Y \\ X_2|Y \\ \vdots \\ X_n|Y \end{pmatrix} = \begin{pmatrix} \mathbb{E}(X_1|Y) \\ \mathbb{E}(X_2|Y) \\ \vdots \\ \mathbb{E}(X_n|Y) \end{pmatrix} = \begin{pmatrix} \mathbb{E}X_1 + \Sigma_{XY,1}\Sigma_{YY}^{-1}(Y - \mu_Y) \\ \mathbb{E}X_2 + \Sigma_{XY,2}\Sigma_{YY}^{-1}(Y - \mu_Y) \\ \vdots \\ \mathbb{E}X_n + \Sigma_{XY,n}\Sigma_{YY}^{-1}(Y - \mu_Y) \end{pmatrix} \\ &= \begin{pmatrix} \mathbb{E}X_1 \\ \mathbb{E}X_2 \\ \vdots \\ \mathbb{E}X_n \end{pmatrix} + \begin{pmatrix} \Sigma_{XY,1}\Sigma_{YY}^{-1}(Y - \mu_Y) \\ \Sigma_{XY,2}\Sigma_{YY}^{-1}(Y - \mu_Y) \\ \vdots \\ \Sigma_{XY,n}\Sigma_{YY}^{-1}(Y - \mu_Y) \end{pmatrix} \\ &= \mu_X + \Sigma_{XY}\Sigma_{YY}^{-1}(Y - \mu_Y) \end{aligned}$$

(c) 由 MA(2) 过程  $X_t = \varepsilon_t + \varepsilon_{t-1} + \varepsilon_{t-2}$  可知  $\varepsilon_t \sim N(0, \sigma_\varepsilon^2)$

则  $\mathbb{E}X_t = 0$  且

$$\sigma_\varepsilon^2(k) = \begin{cases} (3-k)\sigma_\varepsilon^2, & k = 0, 1, 2 \\ 0, & k > 2 \end{cases}$$

定义  $X = (X_{t+1}, X_{t+2}, \dots)^\top, Y = (X_{t-1}, X_{t-2}, \dots)^\top$

由 (a) 问可知  $\mathbb{E}_t X_{t+j} = \mathbb{E}[X_{t+j}|X_t, X_{t-1}, X_{t-2}, \dots] = \mathbb{E}X_{t+j} + \Sigma_{XY,j}\Sigma_{YY}^{-1}(Y - \mu_Y) = \Sigma_{XY,j}\Sigma_{YY}^{-1}Y$

$$\text{且 } \Sigma_{YY} = \sigma_\varepsilon^2 \begin{pmatrix} 3 & 2 & 1 & 0 & \cdots & 0 \\ 2 & 3 & 2 & 1 & \cdots & 0 \\ 1 & 2 & 3 & 2 & \cdots & 0 \\ 0 & 1 & 2 & 3 & \cdots & 0 \\ \vdots & \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & 0 & \cdots & 0 \end{pmatrix}, \Sigma_{XY} = \sigma_\varepsilon^2 \begin{pmatrix} 2 & 1 & 0 & 0 & \cdots & 0 \\ 1 & 0 & 0 & 0 & \cdots & 0 \\ 0 & 0 & 0 & 0 & \cdots & 0 \\ 0 & 0 & 0 & 0 & \cdots & 0 \\ \vdots & \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & 0 & \cdots & 0 \end{pmatrix}$$

将  $\Sigma_{YY}, \Sigma_{XY}, Y$  代入即可求得条件期望

$$\mathbb{E}_t X_{t+j} = \begin{cases} \Sigma_{XY,j}\Sigma_{YY}^{-1}Y, & j \leq 2 \\ 0, & j > 2 \end{cases}$$

则当  $j=1$  时,  $X_{t+1}$  与  $X_{t-2}, \dots$  独立

$$\mathbb{E}_t X_{t+1} = \sigma_\varepsilon^2 \begin{pmatrix} 2 & 1 \end{pmatrix} \left[ \sigma_\varepsilon^2 \begin{pmatrix} 3 & 2 \\ 2 & 3 \end{pmatrix} \right]^{-1} \begin{pmatrix} X_t \\ X_{t-1} \end{pmatrix} = \frac{4}{5}X_t - \frac{1}{5}X_{t-1}$$

当  $j=2$  时,  $X_{t+2}$  与  $X_{t-1}, \dots$  独立,  $\mathbb{E}_t X_{t+2} = \sigma_\varepsilon^2 \frac{1}{3\sigma_\varepsilon^2} X_t = \frac{1}{3} X_t$

当  $j>2$  时,  $\mathbb{E}_t X_{t+j} = 0$

2. (a) 又题易得, 对于 AR(2) 过程  $X_t = \phi_1 X_{t-1} + \phi_2 X_{t-2} + \varepsilon_t$

$$\hat{X}_{t+1|t} = \phi_1 X_t + \phi_2 X_{t-1}$$

$$\text{写为矩阵形式可得 } \begin{pmatrix} \hat{X}_{t+1|t} \\ \hat{X}_{t|t} \end{pmatrix} = \begin{pmatrix} \phi_1 & \phi_2 \\ 1 & 0 \end{pmatrix} \begin{pmatrix} X_{t|t} \\ X_{t-1|t} \end{pmatrix} = \begin{pmatrix} \phi_1 & \phi_2 \\ 1 & 0 \end{pmatrix} \begin{pmatrix} X_t \\ X_{t-1} \end{pmatrix}$$

$$\text{进而可得 } \begin{pmatrix} \hat{X}_{t+j|t} \\ \hat{X}_{t+j-1|t} \end{pmatrix} = \begin{pmatrix} \phi_1 & \phi_2 \\ 1 & 0 \end{pmatrix}^j \begin{pmatrix} X_t \\ X_{t-1} \end{pmatrix}$$

$$\text{下求 } \begin{pmatrix} \phi_1 & \phi_2 \\ 1 & 0 \end{pmatrix}^j, \text{ 令 } \Phi = \begin{pmatrix} \phi_1 & \phi_2 \\ 1 & 0 \end{pmatrix}$$

易得  $\Phi$  的两个特征值为  $\lambda_1 = \frac{\phi_1 + \sqrt{\phi_1^2 + 4\phi_2}}{2}, \lambda_2 = \frac{\phi_1 - \sqrt{\phi_1^2 + 4\phi_2}}{2}$  且有  $\lambda_1 + \lambda_2 = \phi_1, -\lambda_1 \lambda_2 = \phi_2$

进而求得  $\Phi$  的  $P\Lambda P^{-1}$  分解

$$\Phi = \frac{-1}{\phi_2 \sqrt{\phi_1^2 + 4\phi_2}} \begin{pmatrix} \phi_2 & \phi_2 \\ \frac{\sqrt{\phi_1^2 + 4\phi_2} - \phi_1}{2} & \frac{-\sqrt{\phi_1^2 + 4\phi_2} - \phi_1}{2} \end{pmatrix} \begin{pmatrix} \frac{\sqrt{\phi_1^2 + 4\phi_2} + \phi_1}{2} & 0 \\ 0 & \frac{\phi_1 - \sqrt{\phi_1^2 + 4\phi_2}}{2} \end{pmatrix} \begin{pmatrix} \frac{-\sqrt{\phi_1^2 + 4\phi_2} - \phi_1}{2} & -\phi_2 \\ \frac{\sqrt{\phi_1^2 + 4\phi_2} - \phi_1}{2} & \phi_2 \end{pmatrix}$$

$$\text{其中 } P = \begin{pmatrix} \phi_2 & \phi_2 \\ \frac{\sqrt{\phi_1^2 + 4\phi_2} - \phi_1}{2} & \frac{-\sqrt{\phi_1^2 + 4\phi_2} - \phi_1}{2} \end{pmatrix}, \Lambda = \begin{pmatrix} \frac{\sqrt{\phi_1^2 + 4\phi_2} + \phi_1}{2} & 0 \\ 0 & \frac{\phi_1 - \sqrt{\phi_1^2 + 4\phi_2}}{2} \end{pmatrix},$$

$$P^{-1} = \frac{-1}{\phi_2 \sqrt{\phi_1^2 + 4\phi_2}} \begin{pmatrix} \frac{-\sqrt{\phi_1^2 + 4\phi_2} - \phi_1}{2} & -\phi_2 \\ \frac{\sqrt{\phi_1^2 + 4\phi_2} - \phi_1}{2} & \phi_2 \end{pmatrix}$$

$$\text{进而 } \Phi^j = P\Lambda^j P^{-1} = \frac{1}{\lambda_2 - \lambda_1} \begin{pmatrix} -\lambda_1^{j+1} + \lambda_2^{j+1} & \lambda_1^{j+1} \lambda_2 - \lambda_1 \lambda_2^{j+1} \\ -\lambda_1^j + \lambda_2^j & \lambda_1^j \lambda_2 - \lambda_1 \lambda_2^j \end{pmatrix} \text{ (省略计算过程)}$$

$$\text{因此, 可得 } \hat{X}_{t+j|t} = \frac{-\lambda_1^{j+1} + \lambda_2^{j+1}}{\lambda_2 - \lambda_1} X_t + \frac{\lambda_1^{j+1} \lambda_2 - \lambda_1 \lambda_2^{j+1}}{\lambda_2 - \lambda_1} X_{t-1}$$

(b) 见 (a)

$$(c) \text{ 令 } A = \begin{pmatrix} \phi_1 & \phi_2 \\ 1 & 0 \end{pmatrix}$$

可知其特征多项式为  $\lambda^2 - \theta_1 \lambda - \theta_2 = 0$ , 且根为  $\lambda_1 = \frac{1}{z_1}, \lambda_2 = \frac{1}{z_2}$

由于  $X_t$  是平稳时间序列, 则根位于单位圆之外, 即  $|\lambda_1|$  和  $|\lambda_2|$  均小于 1

$$\text{将 } A \text{ 特征分解可得 } A = C\Lambda C^{-1} \text{ 且 } \Lambda = \begin{pmatrix} \lambda_1 & \\ & \lambda_2 \end{pmatrix}$$

$$\text{则 } A^j = (C\Lambda C^{-1})^j = C\Lambda^j C^{-1}$$

$$\text{则 } \lim_{j \rightarrow \infty} A^j = \lim_{j \rightarrow \infty} C \begin{pmatrix} \lambda_1^j & \\ & \lambda_2^j \end{pmatrix} C^{-1} = \begin{pmatrix} 0 & 0 \\ 0 & 0 \end{pmatrix}$$

此时有  $\lim_{j \rightarrow \infty} \begin{pmatrix} \hat{X}_{t+j|t} \\ \hat{X}_{t+j-1|t} \end{pmatrix} = \lim_{j \rightarrow \infty} \begin{pmatrix} \phi_1 & \phi_2 \\ 1 & 0 \end{pmatrix}^j \begin{pmatrix} X_t \\ X_{t-1} \end{pmatrix} = \begin{pmatrix} 0 & 0 \\ 0 & 0 \end{pmatrix}$

则  $\lim_{j \rightarrow \infty} \hat{X}_{t+j|t} = 0$

(d) 当  $\phi_1 = 1.7, \phi_2 = -0.72$  时,  $A = \begin{pmatrix} 1.7 & -0.72 \\ 1 & 0 \end{pmatrix}$

此时 A 的特征多项式为  $\lambda^2 - 1.7\lambda + 0.72 = 0$ , 且根为  $\lambda_1 = 0.8, \lambda_2 = 0.9$

由此可求出  $\lambda_1, \lambda_2$  的特征向量分别为  $T_1 = \begin{pmatrix} 0.8 \\ 1 \end{pmatrix}, T_2 = \begin{pmatrix} 0.8 \\ 1 \end{pmatrix}$

则  $C = \begin{pmatrix} 0.8 & 0.9 \\ 1 & 1 \end{pmatrix}, \Lambda = \begin{pmatrix} 0.8 & \\ & 0.9 \end{pmatrix}$

则  $A = C\Lambda C^{-1} = \begin{pmatrix} 0.8 & 0.9 \\ 1 & 1 \end{pmatrix} \begin{pmatrix} 0.8 & \\ & 0.9 \end{pmatrix} \begin{pmatrix} 0.8 & 0.9 \\ 1 & 1 \end{pmatrix}^{-1}$

且

$$\begin{aligned} \begin{pmatrix} \hat{X}_{t+j|t} \\ \hat{X}_{t+j-1|t} \end{pmatrix} &= A^j \begin{pmatrix} X_t \\ X_{t-1} \end{pmatrix} \\ &= C\Lambda^j C^{-1} \begin{pmatrix} X_t \\ X_{t-1} \end{pmatrix} \\ &= \begin{pmatrix} 0.8 & 0.9 \\ 1 & 1 \end{pmatrix} \begin{pmatrix} 0.8^j & \\ & 0.9^j \end{pmatrix} \begin{pmatrix} 0.8 & 0.9 \\ 1 & 1 \end{pmatrix}^{-1} \begin{pmatrix} X_t \\ X_{t-1} \end{pmatrix} \end{aligned}$$

故  $\hat{X}_{t+j|t} = 10(0.9^{j+1} - 0.8^{j+1})X_t + (9 \cdot 0.8^{j+1} - 8 \cdot 0.9^{j+1})X_{t-1}$