

2020 秋季本科时间序列

第 6 次作业答案

12 月 10 日

1. (a) 由模型 $Y = X\beta + e$, 其拟合值为 $\hat{Y} = X + \hat{\beta}$, 在 OLS 估计下, $\hat{\beta} = (X^T X)^{-1} X^T Y$, 可得

$$\hat{e} = Y - \hat{Y} = [I - X(X^T X)^{-1} X^T]Y = [I - X(X^T X)^{-1} X^T][X\beta + e] = [I - X(X^T X)^{-1} X^T]e$$

- (b) 设 $A = [a_{ij}]_{n \times k}$, $B = [b_{ij}]_{k \times n}$, 同时令 $C = AB = [c_{ij}]_{n \times m}$, $D = BA = [d_{ij}]_{k \times k}$. 对于矩阵 C 上的第 m ($1 \leq m \leq n$) 个对角元素 C_{mm} , 易得

$$C_{mm} = \sum_{j=1}^k a_{mj} b_{jm}$$

进而 $\text{tr}(C) = \text{tr}(AB) = \sum_{i=1}^n \sum_{j=1}^k a_{ij} b_{ji}$. 同理, 对于矩阵 D , 可得 $\text{tr}(D) = \text{tr}(BA) = \sum_{i=1}^k \sum_{j=1}^n b_{ij} a_{ji} = \sum_{j=1}^n \sum_{i=1}^k a_{ji} b_{ij}$, 显然有 $\sum_{i=1}^n \sum_{j=1}^k b_{ij} a_{ji} = \sum_{j=1}^n \sum_{i=1}^k a_{ji} b_{ij}$ (交换积分次序).

故 $\text{tr}(C) = \text{tr}(AB) = \text{tr}(BA) = \text{tr}(D)$. 若 A, B 均为 $n \times n$ 的矩阵, 则

$$\text{tr}(A - B) = \sum_{i=1}^n (a_{ii} - b_{ii}) = \sum_{i=1}^n a_{ii} - \sum_{i=1}^n b_{ii} = \text{tr}(A) - \text{tr}(B)$$

- (c) 由 $\hat{e} = [\hat{e}_1, \hat{e}_2, \dots, \hat{e}_t]^T$, 故 $\sum_t \hat{e}_t^2 = (\hat{e}^T \hat{e})_{1 \times 1} = \text{tr}(\hat{e}^T \hat{e})$. 又由 (a) 有 $\hat{e} = [I - X(X^T X)^{-1} X^T]e$, 故可得

$$\begin{aligned} \hat{e}^T \hat{e} &= [[I - X(X^T X)^{-1} X^T]e]^T [[I - X(X^T X)^{-1} X^T]e] \\ &= e^T [I - X(X^T X)^{-1} X^T]^T [I - X(X^T X)^{-1} X^T]e \\ &= e^T [I - 2X(X^T X)^{-1} X^T + X(X^T X)^{-1} X^T]e \\ &= e^T [I - X(X^T X)^{-1} X^T]e \end{aligned}$$

故 $\text{tr}(\hat{e}^T \hat{e}) = \text{tr}(e^T [I - X(X^T X)^{-1} X^T]e)$, 综上得证 $\sum_t \hat{e}_t^2 = \text{tr}(\hat{e}^T \hat{e}) = \text{tr}(e^T [I - X(X^T X)^{-1} X^T]e)$.

(d) 由 (a),(b),(c) 问结论可知

$$\begin{aligned}
 \mathbb{E}\left[\sum_t \hat{\varepsilon}_t^2 | X\right] &= \mathbb{E}[\text{tr}(\hat{\varepsilon}^T \hat{\varepsilon}) | X] \\
 &= \mathbb{E}[\text{tr}(e^T [I - X(X^T X)^{-1} X^T] e) | X] \\
 &= \mathbb{E}[\text{tr}([I - X(X^T X)^{-1} X^T] e e^T) | X] \\
 &= \mathbb{E}[\text{tr}(e e^T) | X] - \mathbb{E}[\text{tr}(X(X^T X)^{-1} X^T e e^T) | X]
 \end{aligned}$$

又因为 $\mathbb{E}[\text{tr}(A)] = \mathbb{E}[a_{11} + a_{22} + \cdots + a_{nn}] = \mathbb{E}(a_{11}) + \mathbb{E}(a_{22}) + \cdots + \mathbb{E}(a_{nn}) = \text{tr}(\mathbb{E}[A])$ 。
所以上式进一步简化，可得

$$\begin{aligned}
 &\mathbb{E}[\text{tr}(e e^T) | X] - \mathbb{E}[\text{tr}(X(X^T X)^{-1} X^T e e^T) | X] \\
 &= \text{tr}(\mathbb{E}[\text{tr}(e e^T) | X]) - \text{tr}(\mathbb{E}[\text{tr}(X(X^T X)^{-1} X^T e e^T) | X]) \\
 &= \text{tr}(\mathbb{E}[\text{tr}(e e^T) | X]) - \text{tr}(\mathbb{E}[\text{tr}(X(X^T X)^{-1} X^T \mathbb{E}(e e^T | X))]) \\
 &= \text{tr}(\sigma_\varepsilon^2 I_{T \times T}) - \text{tr}((X^T X)^{-1} X^T \sigma_\varepsilon^2) \\
 &= T \sigma_\varepsilon^2 - \text{tr}(I_{k \times k} \sigma_\varepsilon^2) \\
 &= T \sigma_\varepsilon^2 - K \sigma_\varepsilon^2 = (T - K) \sigma_\varepsilon^2
 \end{aligned}$$

由于

$$\mathbb{E}\left(\frac{1}{T-K} \sum_t \hat{\varepsilon}_t^2\right) = \frac{1}{T-K} \mathbb{E}[\mathbb{E}(\sum_t \hat{\varepsilon}_t^2 | X)] = \frac{1}{T-K} \mathbb{E}[(T-K)\sigma_\varepsilon^2] = \sigma_\varepsilon^2$$

故 $\frac{1}{T-K} \sum_t \hat{\varepsilon}_t^2$ 是 σ_ε^2 的无偏估计量。

2. (a) 对于任意 $k \in N$ 且 $k < T$ ，可计算 MA(2) 过程 X_t 序列的自协方差为

$$\sigma^2(k) = \text{cov}(X_t, X_{t-k}) = \mathbb{E}X_t X_{t-k} - 0 = \begin{cases} (1 + \theta_1^2 + \theta_2^2)\sigma_\varepsilon^2, & k = 0 \\ (\theta_1 + \theta_1\theta_2)\sigma_\varepsilon^2, & k = 1 \\ \theta_2\sigma_\varepsilon^2, & k = 2 \\ 0, & k \geq 3 \end{cases}$$

所以 $X = [X_1, \cdots, X_T]^T$ 的协方差矩阵为一个五对角矩阵，如下：

$$\Sigma = \begin{bmatrix} (1 + \theta_1^2 + \theta_2^2)\sigma_\varepsilon^2 & (\theta_1 + \theta_1\theta_2)\sigma_\varepsilon^2 & \theta_2\sigma_\varepsilon^2 & \cdots & 0 \\ (\theta_1 + \theta_1\theta_2)\sigma_\varepsilon^2 & (1 + \theta_1^2 + \theta_2^2)\sigma_\varepsilon^2 & (\theta_1 + \theta_1\theta_2)\sigma_\varepsilon^2 & \cdots & \cdot \\ \theta_2\sigma_\varepsilon^2 & (\theta_1 + \theta_1\theta_2)\sigma_\varepsilon^2 & (1 + \theta_1^2 + \theta_2^2)\sigma_\varepsilon^2 & \cdots & \cdot \\ \vdots & \vdots & \ddots & \vdots & \vdots \\ 0 & \cdots & \cdots & \cdots & (1 + \theta_1^2 + \theta_2^2)\sigma_\varepsilon^2 \end{bmatrix}_{T \times T}$$

可写出对数似然函数

$$\log L(\theta_1, \theta_2, \sigma_\varepsilon^2 | X) = -\frac{1}{2} \log \det \Sigma - \frac{1}{2} X^T \Sigma^{-1} X$$

理论上，通过数值方法可以求解出 $\det \Sigma$ 和 Σ^{-1} 。

(b) 首先

$$|\Sigma| = \begin{vmatrix} (1 + \theta_1^2 + \theta_2^2)\sigma_\varepsilon^2 & (\theta_1 + \theta_1\theta_2)\sigma_\varepsilon^2 & \theta_2\sigma_\varepsilon^2 & \cdots & 0 \\ (\theta_1 + \theta_1\theta_2)\sigma_\varepsilon^2 & (1 + \theta_1^2 + \theta_2^2)\sigma_\varepsilon^2 & (\theta_1 + \theta_1\theta_2)\sigma_\varepsilon^2 & \cdots & \cdot \\ \theta_2\sigma_\varepsilon^2 & (\theta_1 + \theta_1\theta_2)\sigma_\varepsilon^2 & (1 + \theta_1^2 + \theta_2^2)\sigma_\varepsilon^2 & \cdots & \cdot \\ \vdots & \vdots & \ddots & \vdots & \vdots \\ 0 & \cdots & \cdots & \cdots & (1 + \theta_1^2 + \theta_2^2)\sigma_\varepsilon^2 \end{vmatrix}$$

$$= \sigma_\varepsilon^{2T} \begin{vmatrix} 1 + \theta_1^2 + \theta_2^2 & \theta_1 + \theta_1\theta_2 & \theta_2 & \cdots & 0 \\ \theta_1 + \theta_1\theta_2 & 1 + \theta_1^2 + \theta_2^2 & \theta_1 + \theta_1\theta_2 & \cdots & \cdot \\ \theta_2 & \theta_1 + \theta_1\theta_2 & 1 + \theta_1^2 + \theta_2^2 & \cdots & \cdot \\ \vdots & \vdots & \ddots & \vdots & \vdots \\ 0 & \cdots & \cdots & \cdots & 1 + \theta_1^2 + \theta_2^2 \end{vmatrix}$$

令 $\log L$ 分别对 $\theta_1, \theta_2, \sigma_\varepsilon^2$ 求偏导，得

$$\begin{cases} \frac{\partial \log L}{\partial \theta_1} = -\frac{1}{2 \det \Sigma} * \frac{\partial \det \Sigma}{\partial \theta_1} - \frac{1}{2} * \frac{\partial X^T \Sigma^{-1} X}{\partial \theta_1} = 0 \\ \frac{\partial \log L}{\partial \theta_2} = -\frac{1}{2 \det \Sigma} * \frac{\partial \det \Sigma}{\partial \theta_2} - \frac{1}{2} * \frac{\partial X^T \Sigma^{-1} X}{\partial \theta_2} = 0 \\ \frac{\partial \log L}{\partial \sigma_\varepsilon^2} = -\frac{T}{2} * \frac{1}{\sigma_\varepsilon^2} - \frac{1}{2} * \frac{\partial X^T \Sigma^{-1} X}{\partial \sigma_\varepsilon^2} = 0 \end{cases}$$

利用数值解法可解出 $\theta_1, \theta_2, \sigma_\varepsilon^2$ 。

当 $\theta_2 = 0$ 时，

$$\Sigma = \sigma_\varepsilon^2 \begin{bmatrix} (1 + \theta_1^2) & \theta_1 & 0 & \cdots & 0 \\ \theta_1 & (1 + \theta_1^2) & \theta_1 & \cdots & \cdot \\ 0 & \theta_1 & (1 + \theta_1^2) & \cdots & \cdot \\ \vdots & \vdots & \ddots & \vdots & \vdots \\ 0 & \cdots & \cdots & \cdots & (1 + \theta_1^2) \end{bmatrix}$$

利用三对角矩阵行列式公式可得 $\det \Sigma = \frac{1 - \theta_1^{2(T+1)}}{1 - \theta_1^2} \sigma_\varepsilon^2$ 。同理，此处 Σ^{-1} 也可由数值方法求解。进而解法与上文同理。

(c) 由题意及 (a) 问可得方程组

$$\sigma^2(0) = (1 + \theta_1^2 + \theta_2^2)\sigma_\varepsilon^2 = a_0 \quad (1)$$

$$\sigma^2(1) = (\theta_1 + \theta_1\theta_2)\sigma_\varepsilon^2 = a_1 \quad (2)$$

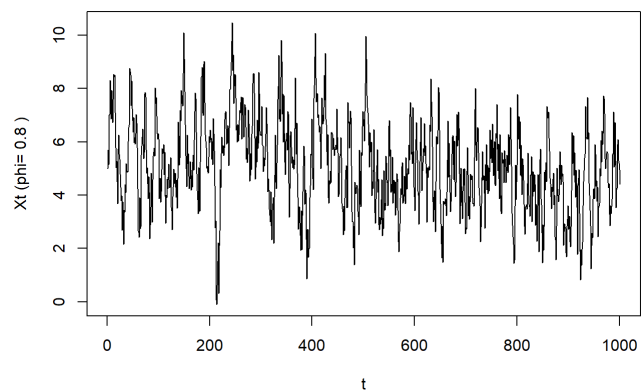
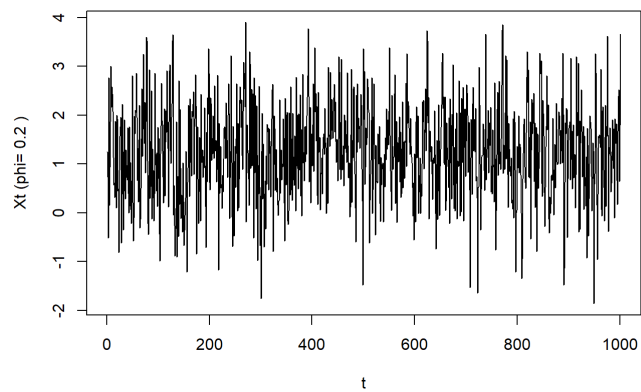
$$\sigma^2(2) = \theta_2\sigma_\varepsilon^2 = a_2 \quad (3)$$

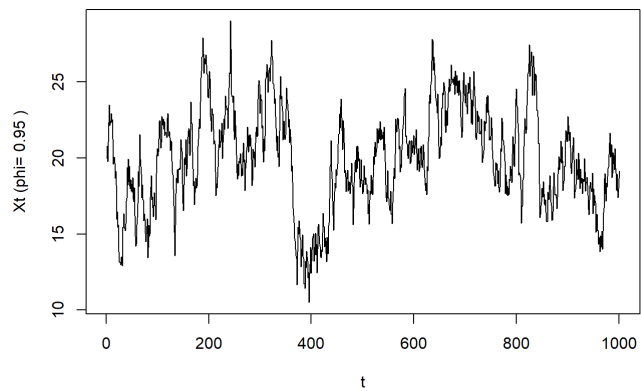
由 (3) 式可得 $\theta_2 = \frac{a_2}{\sigma_\varepsilon^2}$ ，代入 (2) 式，得 $\theta_1(1 + \frac{a_2}{\sigma_\varepsilon^2})\frac{a_2}{\sigma_\varepsilon^2} = a_1$ ，进而 $\theta_1 = \frac{a_1}{\sigma_\varepsilon^2 + a_1}$ 。

将 θ_1 和 θ_2 中关于 σ_ε^2 的表达式代入 (1) 式，得 $(1 + \frac{a_1^2}{(\sigma_\varepsilon^2 + a_1)^2} + \frac{a_2^2}{(\sigma_\varepsilon^2)^2})\sigma_\varepsilon^2 = a_0$ ，理论上可解出 σ_ε^2 （四次方程）。进而代回 θ_1 和 θ_2 关于 σ_ε^2 的式子，可解出 θ_1 和 θ_2 。

3. (a) 代码如下:

```
1 library(tidyverse)
2 mu <- 1
3 phi <- c(0.2, 0.8, 0.95)
4 X <- tibble()
5 for(i in 1:3){
6   X[1, i] = mu/(1 - phi[i])
7   for(j in 1:1000){
8     X[j+1, i] = mu + phi[i]*X[j, i] + rnorm(1,0,1)
9   }
10  plot(X[[i]], type = "l", xlab = "t", ylab = paste('Xt (phi
11         =', phi[i],')'))
}
```





随着 ϕ 的增大，序列的波动性变大，但都没有呈现出明显的趋势及季节性，都是平稳时间序列。

(b) 代码如下：

```

1 beta0 <- tibble()
2 beta1 <- tibble()
3 for(i in 1:3){
4   for(k in 1:10){
5     y <- unlist(X[2:(100*k+1), i])
6     x <- unlist(X[1:(100*k), i])
7     ols <- coefficients(lm(y ~ x))
8     beta0[k, i] = ols[1]
9     beta1[k, i] = ols[2]
10  }
11  plot(beta0[[i]], type = "l", xlab = "k", ylab = paste('
      beta0 (phi=', phi[i], ')'))
12  plot(beta1[[i]], type = "l", xlab = "k", ylab = paste('
      beta1 (phi=', phi[i], ')'))
13 }

```

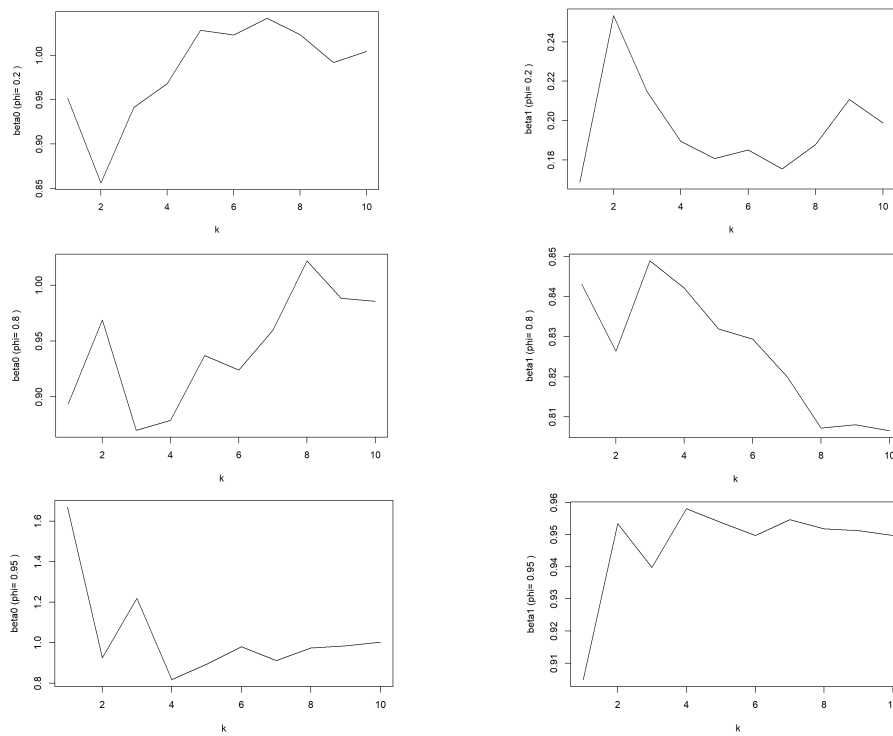
随着 ϕ 的增大，OLS 估计随样本量在 100k 增加时的收敛性逐渐增强。

(c) 代码如下：

```

1 # phi = 0.2
2 X1 <- tibble()
3 for(i in 1:500){
4   X1[1, i] = mu/(1 - phi[1])
5   for(j in 1:100){
6     X1[j+1, i] = mu + phi[1]*X1[j, i] + rnorm(1,0,1)

```



```

7   }
8   }
9   # phi = 0.8
10  X2 <- tibble()
11  for(i in 1:500){
12    X2[1, i] = mu/(1 - phi[2])
13    for(j in 1:100){
14      X2[j+1, i] = mu + phi[2]*X2[j, i] + rnorm(1,0,1)
15    }
16  }
17  # phi = 0.95
18  X3 <- tibble()
19  for(i in 1:500){
20    X3[1, i] = mu/(1 - phi[3])
21    for(j in 1:100){
22      X3[j+1, i] = mu + phi[3]*X3[j, i] + rnorm(1,0,1)
23    }
24  }
25  X <- list(X1, X2, X3)

```

```

26
27 beta0 <- tibble()
28 beta1 <- tibble()
29 sd_beta0 <- vector("double",3)
30 sd_beta1 <- vector("double",3)
31 for(i in 1:3){
32   for(j in 1:500){
33     y <- unlist(X[[i]][2:101, j])
34     x <- unlist(X[[i]][1:100, j])
35     ols <- coefficients(lm(y ~ x))
36     beta0[j, i] <- ols[1]
37     beta1[j, i] <- ols[2]
38   }
39   hist(beta0[[i]], xlab = paste('beta0 (phi=', phi[i],')'),
40         main = paste('Histogram of beta0 (phi=', phi[i],')'))
41   hist(beta1[[i]], xlab = paste('beta1 (phi=', phi[i],')'),
42         main = paste('Histogram of beta1 (phi=', phi[i],')'))
43   sd_beta0[i] <- sd(beta0[[i]])
44   sd_beta1[i] <- sd(beta1[[i]])
45 }
46 sd_beta0
47 ## [1] 0.1581694 0.3707317 0.9623762
48 sd_beta1
49 ## [1] 0.09468137 0.06826351 0.04768194

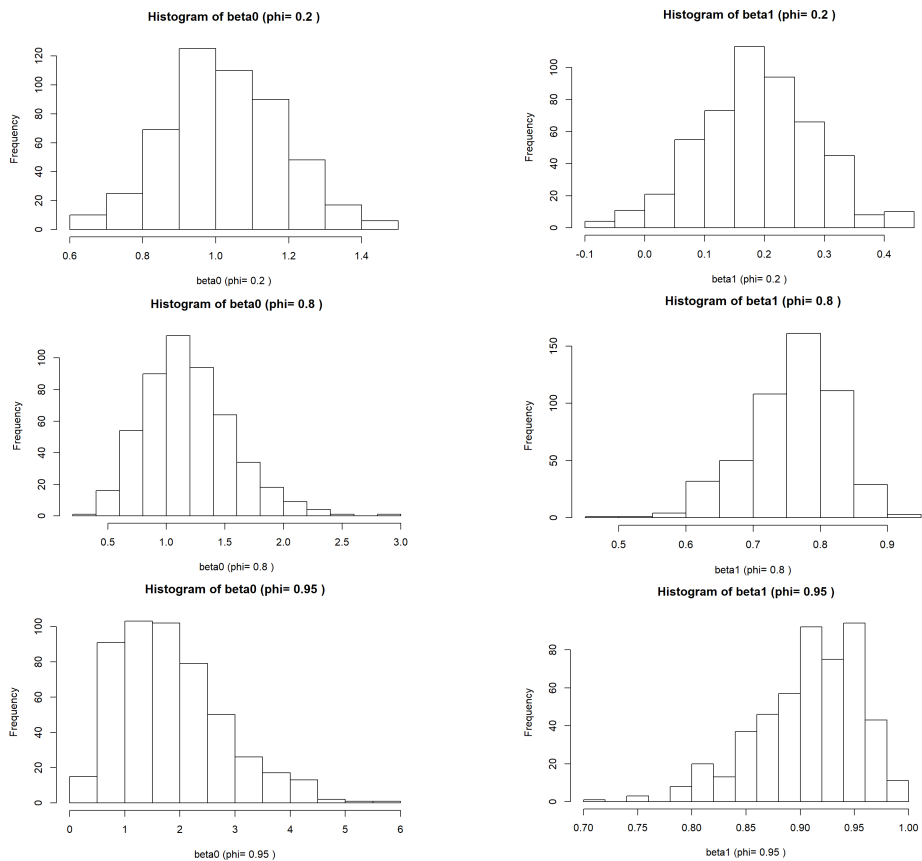
```

(d) 代码如下:

```

1   x <- vector("double",101)
2   sd_beta0 <- vector("double",3)
3   sd_beta1 <- vector("double",3)
4   for (i in 1:3) {
5     x[1] <- mu/(1 - phi[i])
6     for (j in 1:100) {
7       x[j+1] <- mu + phi[i]*x[j] + rnorm(1,0,1)
8     }
9     X <- cbind(rep(1,100), x[2:101])
10    M <- t(X) %*% X / 100 # 求M矩阵

```



```

11 sd_beta0[i] <- sqrt(1/100*1*solve(M)[1,1])
12 sd_beta1[i] <- sqrt(1/100*1*solve(M)[2,2])
13 }
14 sd_beta0
15 ## [1] 0.1515409 0.3606424 1.0093074
16 sd_beta1
17 ## [1] 0.09815747 0.07496892 0.04951775

```

渐进标准误接近样本标准差的值。

(e) 代码如下:

```

1 # phi = 0.2
2 X1 <- tibble()
3 for(i in 1:500){
4   X1[1, i] = mu/(1 - phi[1])
5   for(j in 1:900){
6     X1[j+1, i] = mu + phi[1]*X1[j, i] + rnorm(1,0,1)
7   }
8 }

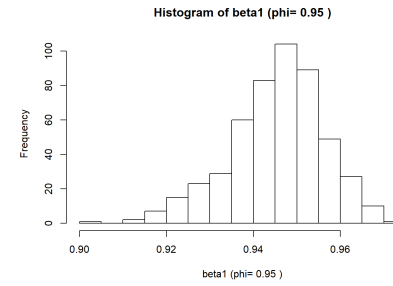
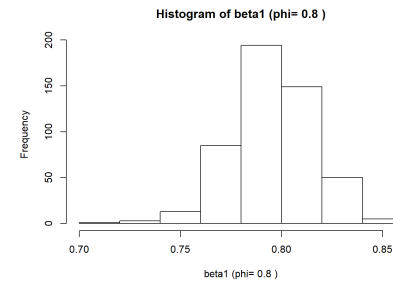
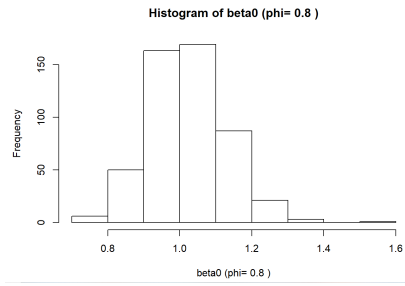
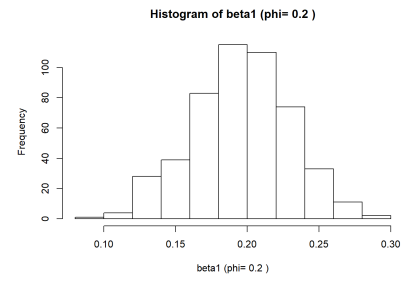
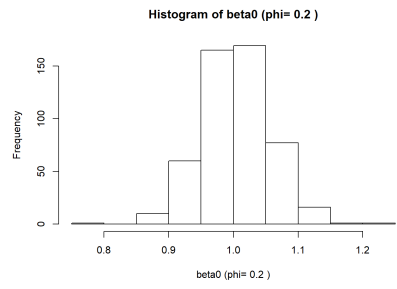
```



```

9 # phi = 0.8
10 X2 <- tibble()
11 for(i in 1:500){
12   X2[1, i] = mu/(1 - phi[2])
13   for(j in 1:900){
14     X2[j+1, i] = mu + phi[2]*X2[j, i] + rnorm(1,0,1)
15   }
16 }
17 # phi = 0.95
18 X3 <- tibble()
19 for(i in 1:500){
20   X3[1, i] = mu/(1 - phi[3])
21   for(j in 1:900){
22     X3[j+1, i] = mu + phi[3]*X3[j, i] + rnorm(1,0,1)
23   }
24 }
25 X <- list(X1, X2, X3)
26
27 beta0 <- tibble()
28 beta1 <- tibble()
29 sd_beta0 <- vector("double",3)
30 sd_beta1 <- vector("double",3)
31 for(i in 1:3){
32   for(j in 1:500){
33     y <- unlist(X[[i]][2:901, j])
34     x <- unlist(X[[i]][1:900, j])
35     ols <- coefficients(lm(y ~ x))
36     beta0[j, i] <- ols[1]
37     beta1[j, i] <- ols[2]
38   }
39   hist(beta0[[i]], xlab = paste('beta0 (phi=', phi[i],')'),
40         main = paste('Histogram of beta0 (phi=', phi[i],')'))
41   hist(beta1[[i]], xlab = paste('beta1 (phi=', phi[i],')'),
42         main = paste('Histogram of beta1 (phi=', phi[i],')'))
43   sd_beta0[i] <- sd(beta0[[i]])

```



```
42 sd_beta1[i] <- sd(beta1[[i]])
43 }
```

此时估计系数的样本标准差近似为 (c) 中的 $1/3$ 。